



Calhoun: The NPS Institutional Archive

DSpace Repository

Theses and Dissertations

1. Thesis and Dissertation Collection, all items

1971-03

An approximation to offset circle probabilities.

Wrockloff, George Edmund

Monterey, California ; Naval Postgraduate School

https://hdl.handle.net/10945/15584

Downloaded from NPS Archive: Calhoun



Calhoun is the Naval Postgraduate School's public access digital repository for research materials and institutional publications created by the NPS community. Calhoun is named for Professor of Mathematics Guy K. Calhoun, NPS's first appointed -- and published -- scholarly author.

> Dudley Knox Library / Naval Postgraduate School 411 Dyer Road / 1 University Circle Monterey, California USA 93943

http://www.nps.edu/library

AN APPROXIMATION TO OFFSET CIRCLE PROBABILITIES

By

George Edmund Wrockloff



United States Naval Postgraduate School				
THESIS				
AN APPROXIMATION TO OFFSET CIRCLE PROBABILITIES				
by				
George Edmund Wrockloff,III				
Thesis Advisor: J. G. Taylor				
March 1971				

Approved for public release; distribution unlimited.

4 . L .

.

An Approximation to Offset Circle Probabilities

by

George Edmund Wrockloff,III Major, United States Army B. S., United States Military Academy, 1960

Submitted in partial fulfillment of the requirements for the degree of

MASTER OF SCIENCE IN OPERATIONS RESEARCH

from the

NAVAL POSTGRADUATE SCHOOL March 1971



NAVAL FOR GRADUALE SCHOOL MONTEREY, CALIF. 95940

ABSTRACT

If a weapon system is fired at a circular target and the impact point is distributed as a bivariate normal random variable, it is not particularly difficult to determine hit probability if the expected impact point is at target center. If, for some reason, the expected impact point is offset, the problem of determining hit probabilities becomes quite complex. Herein is developed a method to approximate such a hit probability.

TABLE OF CONTENTS

.

I.	INTRODUCTION	1
II.	NATURE OF THE PROBLEM	6
III.	CONCLUSION	12
IV.	SUMMARY	17
APPENI	DIX A: Comparison of Hit Probabilities	18
BIBLI	OGRAPHY	19
INITI	AL DISTRIBUTION LIST	20
FORM I	DD 1473	21

c

I. INTRODUCTION

The probability of hitting a specified target with a given weapon system is of particular interest to both military commanders and weapon systems analysts. Determining hit probability is not always a simple matter. In particular, if the impact points of the weapon system are distributed according to some probability distribution, and the expected impact point is not at target center, the problem of determining hit probabilities becomes quite complex.

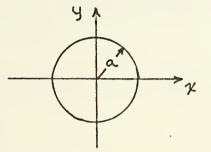
Over the past thirty years this problem has been studied in depth as the extensive bibliographies of F. Grubbs [Ref. 1], A. Eckler [Ref. 2], and W. Guenther and P. Terragno [Ref. 3] illustrate. Several approaches to solving the problem have been investigated by these and other authors in an effort to simplify the computations required to estimate or actually compute the hit probabilities.

As an example of one approach to the problem, Grubbs [Ref. 1] uses what he calls a "single, straight-forward and rather simple technique" to approximate these hit probabilities in which he employs an approximate central chi-square distribution with fractional degree of freedom, or a transformation to approximate normality. There are, of course, other methods available to approximate these hit probabilities, one of which will be discussed in the following pages.

The method which will be discussed is computationally easy, using simple mathematics in conjunction with a table of modified Bessel functions. These tables are usually available to the analyst and enable him to accurately approximate the desired hit probabilities without the use of complex computations. It should be mentioned that tables of offset circle probabilities have been compiled by Rand Corporation [Ref. 4]. These tables give hit probabilities as a function of target radius and offset distance of expected impact point, but are not as readily available as the Bessel function tables used in the following procedure. It should also be noted that routines for Bessel functions are available for most all computer systems.

II. NATURE OF THE PROBLEM

Consider a circular target with radius a. Place the origin of the coordinate system at the center of the target.



Let f(r, y) be the probability density function of the impact point, (x, y), of the projectile fired at the target. Let (μ_x, μ_y) be the mean of the impact point distribution and σ_x^2 and σ_y^2 be the variance of the impact point distribution in the x and y directions respectively. Assume covariance is zero. If the distribution of impact points follows a bivariate normal distribution, then

$$f(x,y) = \frac{1}{2\pi\sigma_x\sigma_y} \exp\left\{-\frac{1}{2}\left[\left(\frac{x-\mu_v}{\sigma_x}\right)^2 + \left(\frac{y-\mu_y}{\sigma_y}\right)^2\right]\right\}$$

Assume $\sigma_x = \sigma_y = \sigma$, and $R^2 = \mu_x^2 + \mu_y^2$. Let P_h be the probability of a hit on the target, then

$$P_{h} = \frac{1}{2\pi\sigma^{2}} \int \int exp \left\{ \frac{-1}{2\sigma^{2}} \left[(x - \mu_{k})^{2} + (y - \mu_{y})^{2} \right] \right\} dx dy$$

$$\chi^{2} \cdot y^{2} \in a^{2}$$

Changing to polar coordinates yields

$$\chi = r \cos \theta$$
, $y = r \sin \theta$

and

$$P_{h} = \frac{1}{2\pi\sigma^{2}} \int_{0}^{2\pi} d\theta \int_{0}^{r} exp\left\{-\frac{1}{2\sigma^{2}}\left[\left(r\cos\theta - M_{x}\right)^{2} + \left(r\sin\theta - M_{y}\right)^{2}\right]\right\} dr$$
$$= \frac{e^{2\pi\sigma^{2}}}{2\pi\sigma^{2}} \int_{0}^{2\pi} d\theta \int_{0}^{r} exp\left\{-\frac{1}{2\sigma^{2}}\left[r^{2} - 2r\left(\cos\theta M_{x} + \sin\theta M_{y}\right)\right]\right\} dr$$
$$= \frac{e^{2\pi\sigma^{2}}}{2\pi\sigma^{2}} \int_{0}^{2\pi} d\theta \int_{0}^{r} exp\left\{-\frac{1}{2\sigma^{2}}\left[r^{2} - 2r\left(\cos\theta M_{y} + \sin\theta M_{y}\right)\right]\right\} dr$$

$$=\frac{e^{z\sigma^{2}}}{z\pi\sigma^{2}}\int_{0}^{r}e_{x}\rho(\frac{-r^{2}}{z\sigma^{2}})\left[\int_{0}^{r}e_{x}\rho\{r_{z}(\cos\theta_{x}+\sin\theta_{y})\}d\theta]dr$$

Rotate axes so that $\mu_{\chi} = R$ and $\mu_{\gamma} = 0$, then

$$P_{h} = \frac{e^{2\sigma^{2}}}{2\pi\sigma^{2}} \int_{0}^{\alpha} e_{x} p\left(\frac{-r^{2}}{2\sigma^{2}}\right) \left[\int_{0}^{2\pi} e_{x} p\left(\frac{rR\cos\theta}{\sigma^{2}}\right) d\theta \right] dr$$

Looking at the term, $\int exp(\frac{rR\cos\theta}{\sigma^2})d\theta$,

$$\sum_{0}^{2\pi} \Theta d\theta = \cos^{n-1} \theta \sin \theta \Big|_{0}^{2\pi} - (n-1) \int_{0}^{2\pi} \cos^{n-2} \theta \sin \theta (-\sin \theta) d\theta$$

$$= (n-1) \int_{0}^{2\pi} \cos^{n-2} \theta \sin^{2} \theta d\theta$$

$$= (n-1) \int_{0}^{2\pi} \cos^{n-2} \theta d\theta - (n-1) \int_{0}^{2\pi} \cos^{n} \theta d\theta$$

and
$$\int_{0}^{2\pi} \cos^{n}\theta d\theta = \left(\frac{n-1}{n}\right)\int_{0}^{2\pi} \cos^{n-2}\theta d\theta$$

If n is odd, then $\int_{0}^{2\pi} \cos^{n}\theta d\theta = 0$. If n is even, then

$$\sum_{n=0}^{2\pi} \int \cos^{n} \theta \, d\theta = \frac{(n-1)(n-3)(n-5)\cdots 3\cdot 1}{n(n-2)(n-4)\cdots 4\cdot 2} Z \pi \qquad \text{Thus,}$$

$$\sum_{n=0}^{2\pi} \int \exp(12\cos\theta) \, d\theta = \sum_{n=0}^{\infty} \frac{\chi^{n}}{n!} \int \cos^{n} \theta \, d\theta = \sum_{n=0}^{\infty} \frac{\chi^{2m}}{(2m)!} \frac{(2m)!}{2^{2m}(m!)^{2}}$$

$$= 2\pi \sum_{m=0}^{\infty} (\frac{\chi}{2})^{2m} \frac{1}{(m!)^{2}} \qquad \text{Also, for n even and equal to } 2m,$$

$$\frac{(n-1)(n-3)(n-5)\cdots 3\cdot 1}{n(n-2)(n-4)\cdots 4\cdot 2} = \frac{n! 2\pi}{[n(n-2)(n-4)\cdots 4\cdot 2]^2} = \frac{n! 2\pi}{[\binom{n}{2}!]^2 2^n}$$

 $= \frac{(2m)! 2\pi}{(m!)^2 z^{2m}}$

The Bessel function, denoted by $J_p(\mathcal{V})$, is equal to $\sum_{n=0}^{\infty} \frac{(-1)^n \left(\frac{\mathcal{V}}{2}\right)^{2n+p}}{n! (n+p)!}$, where x is the argument of the

Bessel function and p is the order of the Bessel function. The modified Bessel function of order zero is denoted by $I_o(r)$, and is equal to $J_o(irc) = \sum_{n=0}^{\infty} \frac{(-i)^n (\frac{irc}{2})^{2n}}{(n!)^2} = \sum_{n=0}^{\infty} \frac{(\frac{r}{2})^{2n}}{(n!)^2}$. Thus, $\int_{0}^{2\pi} erp(rcos\theta) d\theta = 2\pi \sum_{m=0}^{\infty} \frac{(\frac{r}{2})^{2m}}{(m!)^2} = 2\pi I_o(rr)$, so that $\int erp(\frac{rR}{\sigma^2}cos\theta) d\theta = 2\pi I_o(\frac{rR}{\sigma^2})$

It was shown previously that the probability of hitting the target, P_h , was equal to $\frac{R^2}{2\pi\sigma^2} \int revp(\frac{-r^2}{2\sigma^2}) \left[\int evp(\frac{-R\cos\theta}{\sigma^2}) d\theta \right] dr$.

It was also shown that $\int e_{R} e_{T} e_{T} e_{T} e_{T} e_{T} e_{T} \int e_{T} e_{T} e_{T} e_{T} \int e_{T} e_{T}$

It is only necessary to know the ratios, $\frac{R}{\sigma}$ and $\frac{\alpha}{\sigma}$, to determine the probability of hitting the target, $P_{h}(\frac{R}{\sigma},\frac{\alpha}{\sigma}) = \frac{1}{\sigma^{2}} e^{\frac{R^{2}}{2\sigma^{2}}} \int_{\sigma} re^{\frac{r^{2}}{2\sigma^{2}}} I_{o}(\frac{rR}{\sigma^{2}}) dr.$ The Rand Corporation has constructed tables giving thes

The Rand Corporation has constructed tables giving these hit probabilities as a function of the two ratios, $\frac{2}{5}$ and $\frac{2}{5}$, [Ref. 3].

Gilliland, [Ref. 4], shows that

$$P_{h}\left(\frac{R}{\sigma},\frac{\alpha}{\sigma}\right) = e^{\frac{2}{2\sigma^{2}}} \int_{m=0}^{\infty} \frac{\binom{R}{\sigma}^{2m}}{m! 2m} P_{m}\left(\frac{\alpha^{2}}{2\sigma^{2}}\right)$$
where $P_{m}\left(\frac{\alpha^{2}}{2\sigma^{2}}\right) = 1 - e^{\frac{-\alpha^{2}}{2\sigma^{2}}} \int_{n=0}^{\infty} \left(\frac{\alpha^{2}}{2\sigma^{2}}\right)^{n} \frac{1}{n!}$

Luke, [Ref. 5], defines the function

$$J(\gamma, \gamma) = 1 - e^{\gamma} \int_{e^{-t}}^{e^{-t}} I_0[Z(\gamma t)^{\prime z}] dt .$$
Recall that

$$P_1(\stackrel{R}{\sigma}, \stackrel{A}{\sigma}) = \frac{1}{\sigma^2} e^{2\sigma^2} \int_{e^{-\frac{1}{2}\sigma^2}}^{e^{-\frac{1}{2}\sigma^2}} I_0(\frac{rR}{\sigma^2}) dr.$$
It follows that

$$J(\stackrel{A^2}{2\sigma^2}, \stackrel{R^2}{2\sigma^2}) = 1 - e^{-\frac{R^2}{2\sigma^2}} \int_{e^{-t}}^{e^{-t}} I_0[Z(\frac{R}{2\sigma^2}t)^{\prime z}] dt .$$
Let $t = \stackrel{V^2}{2\sigma^2}$, $dt = \frac{vdv}{\sigma^2}$. When $t = \stackrel{A^2}{2\sigma^2}$, $v = a$.

Thus,

$$J(\frac{a^{2}}{2\sigma^{2}},\frac{R^{2}}{2\sigma^{2}}) = 1 - e^{\frac{R^{2}}{2\sigma^{2}}} \int_{\sigma^{2}}^{v} e^{\frac{v^{2}}{2\sigma^{2}}} I_{o} \left[2(\frac{R^{2}}{2\sigma^{2}},\frac{v^{2}}{2\sigma^{2}})^{\prime 2} \right] dv$$
or, $J(\frac{a^{2}}{2\sigma^{2}},\frac{R^{2}}{2\sigma^{2}}) = 1 - \frac{1}{\sigma^{2}} e^{\frac{R^{2}}{2\sigma^{2}}} \int_{\sigma^{2}}^{v} e^{\frac{v^{2}}{2\sigma^{2}}} I_{o} \left(\frac{Rv}{\sigma^{2}}\right) dv$
As a result,
$$P(\frac{R}{\sigma},\frac{a}{\sigma}) = 1 - J(\frac{a^{2}}{2\sigma^{2}},\frac{R^{2}}{2\sigma^{2}}).$$

Luke, [Ref. 5], gives the expansion of $J(\gamma,\gamma)$ in a series of Bessel functions.

$$J(u, y) = e^{-(u+y)} \bigvee_{k=0}^{\infty} \eta^{k_2} I_k(\beta) \quad \text{for } \eta < 1$$

$$J(u, y) = 1 - e^{-(u+y)} \bigvee_{k=1}^{\infty} \eta^{-k} I_k(\beta) \quad \text{for } \eta > 1$$
where $\beta = Z(uy)^{v_2}$, and $\eta = {\binom{y}{\chi}}^{v_2}$.
Consider
$$J(\frac{a^2}{2\sigma^2}, \frac{R^2}{2\sigma^2}) = e^{-(\frac{a^2+R^2}{2\sigma^2})} \bigvee_{k=0}^{\infty} {\binom{R}{a}}^k I_k(\frac{aR}{\sigma^2}) \quad \text{for } R < a$$

$$= 1 - e^{-(\frac{a^2+R^2}{2\sigma^2})} \bigvee_{k=0}^{\infty} {\binom{R}{a}}^k I_k(\frac{aR}{\sigma^2}) \quad \text{for } R > a$$
Then
$$P_h(\frac{R}{\sigma}, \frac{a}{\sigma}) = 1 - e^{-(\frac{a^2+R^2}{2\sigma^2})} \bigvee_{k=0}^{\infty} {\binom{R}{a}}^k I_k(\frac{aR}{\sigma^2}) \quad \text{for } R < a$$

$$= -(\frac{a^2+R^2}{2\sigma^2}) \bigvee_{k=0}^{\infty} {\binom{R}{a}}^k I_k(\frac{aR}{\sigma^2}) \quad \text{for } R < a$$

 $= e^{\frac{2}{2}} \sum_{k=1}^{\infty} \left(\frac{a}{k}\right)^{k} I_{k} \left(\frac{a}{\sigma^{2}}\right) \qquad \text{for } R > a$

where

.

$$R = \int M_{\chi}^{2} + M_{y}^{2}$$
, and $I_{n}(\chi) = \sum_{k=0}^{\infty} \frac{\binom{\chi}{2}^{2k+n}}{k!(n+k)!}$.

These latter two equations yield the true hit probability for a weapon system fired at a circular target with radius a, and normally distributed impact points with expected impact point at μ_X and μ_X .

A good approximation to the true hit probability can be had by considering only the first few terms of the infinite series in the appropriate equation. That is

$$\hat{P}_{h}\left(\frac{R}{\sigma},\frac{a}{\sigma}\right) = 1 - e^{-\left(\frac{a^{2}+R^{2}}{2\sigma^{2}}\right)} \sum_{k=0}^{n} \left(\frac{R}{a}\right)^{k} I_{k}\left(\frac{a}{\sigma^{2}}\right) \qquad R < a$$
$$= e^{-\left(\frac{a^{2}+R^{2}}{2\sigma^{2}}\right)} \sum_{k=1}^{n} \left(\frac{a}{R}\right)^{k} I_{k}\left(\frac{a}{\sigma^{2}}\right) \qquad R > a$$

where n is, hopefully, relatively small. If n is less than five, then the approximation would be beneficial to analysts and would be quite simple to use. If n is greater than five, the calculations required to get a reasonably accurate approximation to hit probability become prohibitive.

III. CONCLUSIONS

The accuracy of the above approximations was checked on an IBM 360 computer. The program used the FORTRAN language and two IBM subroutines, IO and INUE, which computed the modified Bessel functions. This program is available through the Operations Analysis Department of the Naval Postgraduate School in Monterey, California.

The computer program was basically an iterative process in which different values of the ratios $\frac{2}{5}$ and $\frac{2}{5}$ were set and the hit probabilities computed while the number of terms from the appropriate infinite series varied from one to ten. The approximate hit probabilities thus obtained were then compared to the true hit probabilities from the Rand tables [Ref. 4].

The two aspects of primary interest were (1) how accurate are the approximating equations in the various ranges of $\frac{R}{\sigma}$ and $\frac{C}{\sigma}$, and (2) how many terms of the infinite series in the approximating equations must be considered before a reasonable degree of accuracy is obtained. The specific values of the ratios $\frac{R}{\sigma}$ and $\frac{C}{\sigma}$ ranged from .5 to 5.0 in the computer program. For the equation for R<a, the ratio $\frac{R}{\sigma}$ was initialized at the value .5 and the ratio $\frac{C}{\sigma}$ was allowed to vary between .5 and 5.0. The value for $\frac{R}{\sigma}$ was then incremented to 1.0 and the process repeated with values of $\frac{C}{\sigma}$ greater than or equal to 1.0. This procedure was continued until both $\frac{R}{\sigma}$ and $\frac{C}{\sigma}$ were

equal to 5.0. For each value of $\frac{2}{\sigma}$ and $\frac{2}{\sigma}$ the approximate hit probability was computed using only the first term of the infinite series initially, and then again for each additional term used, up to ten terms. In this manner it was possible to see how rapidly the approximating equations converged to the true hit probability.

For the equation for R>a, the same type of procedure was used, holding the ratio $\frac{2}{5}$ constant while the ratio $\frac{2}{5}$ varied through the appropriate values. As before, the ratio $\frac{2}{5}$ was initialized at .5 and then incremented by .5 each time the ratio $\frac{2}{5}$ had run its range of values and the hit probabilities computed.

Representative values of hit probabilities obtained through the approximating equations are contained in Appendix A. Figure 1 compares approximate hit probabilities with actual hit probabilities for R<a, and Fig. 2 does the same for R>a. In both cases the values shown were computed using the first five terms of the infinite series from the appropriate approximating equation.

The convergence of the approximating equation values to the true hit probabilities is shown in the two examples below. Example one shows the convergence of the approximating equation for R<a as the number of terms from the infinite series increases. Values of $\frac{2}{5}$ and $\frac{2}{5}$ were selected to show both rapid convergence and relatively slow convergence. Example two shows convergence of the approximating equation for R>a. These examples along with Appendix A point out that the

approximation of hit probability for R<a is optimistic while the approximation of hit probability for R>a is conservative. Example 1 (R<a)

$$\frac{R}{\sigma} = .5, \frac{a}{\sigma} = 1.0$$

Approximation using:

.2089828 one term two terms .1108707 three terms .1047551 four terms .1045011 five terms .1044937 six terms .1044937 seven terms .1044937 eight terms .1044937 .1044937 nine terms .1044937 ten terms

True probability: .104491

$$\frac{2}{3} = 4.5, \frac{2}{3} = 5.0$$

Approximation using:

one term	.9253550
two terms	.8596848
three terms	.8044760
four terms	.7601165
five terms	.7260436
six terms	.7010158
seven terms	.6834279
eight terms	.6715976
nine terms	.6639764
ten terms	.6592714

True probability: .653154

Example 2 (R>a)

$$\frac{a}{\sigma} = .5, \frac{R}{\sigma} = 1.0$$

Approximation using:

one term	.0690204
two terms	.0732899
three terms	.0734669
four terms	.0734724
five terms	.0734725
six terms	.0734725
seven terms	.0734725
eight terms	.0734725
nine terms	.0734725
ten terms	.0734725

True probability: .073473

$$a = 4.5, R = 5.0$$

Approximation using:

one term	.0656702
two terms	.1208790
three terms	.1652384
four terms	.1993113
five terms	.2243391
six terms	.2419270
seven terms	.2537574
eight terms	.2613786
nine terms	.2660835
ten terms	.2688691

True probability: .272201

Accuracy of the approximating equations can be improved in all cases by considering more terms of the appropriate infinite series. If it is necessary to use more than five terms of the series, however, the mathematic manipulations involved become impractical to do by hand. Accuracy can also

be improved when $\frac{2}{6}$ is equal to $\frac{2}{6}$. From Appendix A it is apparent that when $\frac{2}{6}$ is equal to $\frac{2}{6}$ and both are greater than 3.0, accuracy is the poorest. To more accurately approximate the true hit probability in this case it is sufficient to compute the arithmetic mean of the values obtained from both equations (R<a and R>a). For example, for $\frac{2}{6} = \frac{2}{6} = 5.0$, the true probability is .460. The equation for R<a yields an approximation of .643, while the equation of these two values is .484, which is more accurate than either of the two approximations.

IV. SUMMARY

If the impact point of a weapon system is distributed as a bivariate normal random variable, the task of computing the hit probability for a circular target when the expected impact point is somewhere other than at target center is formidable. The two approximating equations that have been developed

$$\hat{P}_{h}(\hat{\sigma}, \hat{\sigma}) = 1 - e^{-\left(\frac{a^{2} + R^{2}}{2\sigma^{2}}\right)} \sum_{k=0}^{n} \left(\frac{R}{a}\right)^{k} I_{k}\left(\frac{a, R}{\sigma^{2}}\right) \qquad \text{for } R < \alpha$$

$$= e^{-\left(\frac{a^{2} + R^{2}}{2\sigma^{2}}\right)} \sum_{k=1}^{n} \left(\frac{a}{k}\right)^{k} I_{k}\left(\frac{a, R}{\sigma^{2}}\right) \qquad \text{for } R > \alpha$$

are relatively accurate approximations to true hit probabilities using only the first five terms of the series. Greater accuracy can be achieved by considering more terms. To have an accurate approximation to such a hit probability is beneficial to both weapon systems analysts and military commanders, because of the difficulties involved in computing actual hit probabilities. The above equations yield a good approximation to hit probabilities using simple mathematics and a table of modified Bessel functions, which is not difficult to obtain.

17

APPENDIX A

COMPARISON OF HIT PROBABILITIES

R	1.0	2.0	3.0	4.0	5.0
916	app act	2.0 app act	app act	app act	app act
1.0	.267 .267	7 * *	* *	* *	* *
2.0	.731 .731	.410 .397	* *	* *	* *
3.0	.956 .956	.788 .786	.498 .433	* *	* *
4.0	.997 .997	.966 .966	.814 .803	.578 .450	* *
5.0	.999 .999	.998.998	.970 .969	.835 .813	.643 .460

Figure 1. Table for R Less Than a

\mathbf{X}	1.	. 0	2.	0	3.	.0	4.	.0	5.	0
R	app	act	app	act	app	act	app	act	app	act
1.0	.267	.267	*	*	*	*	*	*	5. app *	*
2.0	.082	.082	.393	.397	*	*	*	*	*	*
3.0	.011	.011	.113	.113	.399	.433	*	*	*	*
4.0	.001	.001	.015	.015	.121	.126	.366	.450	*	*
5.0	.000	.000	.001	.001	.016	.017	.120	.133	.325	.460

Figure 2. Table for R Greater Than a

These tables compare the approximation to the actual hit probabilities (app) and the actual hit probabilities (act) themselves. Only the first five terms of the infinite series described on page llwere used for the approximations. Accuracy can be improved in all cases by considering more terms of the series.

18

BIBLIOGRAPHY

- 1. Grubbs, F. E., "Approximate Circular and Noncircular Offset Probabilities of Hitting," <u>Operations Research</u>, v. 12, pp. 51-62, 1964.
- Eckler, A. R., "A Survey of Coverage Proglems Associated with Point and Area Targets," <u>Technometrics</u>, v. II, No. 3, pp. 561-588, August 1969.
- Guenther, W. C. and Terragno, P. J., "A Review of the Literature on a Class of Coverage Problems," Annals of Mathematical Statistics, v. 35, pp. 232-260, 1964.
- 4. Rand Corporation, "Table of Q Functions," Research Memo randum No. 339, 1 January 1950.
- Gilliland, D. C., "Integral of the Bivariate Normal Distribution Over an Offset Circle," Journal of the American Statistical Association, v. 57, No. 300, December 1962.
- 6. Luke, Y., Integrals of Bessel Functions, McGraw-Hill, 1962.

INITIAL DISTRIBUTION LIST

		No. Copies
1.	Defense Documentation Center Cameron Station Alexandria, Virginia 22314	2
2.	Library, Code 0212 Naval Postgraduate School Monterey, California 93940	2
3.	Professor James G. Taylor, Code 55 Ta Department of Operations Analysis Naval Postgraduate School Monterey, California 93940	1
4.	Department of Operations Analysis (Code 55) Naval Postgraduate School Monterey, California 93940	1
5.	Major George E. Wrockloff III, USA (student) 2534 Blossom Drive San Antonio, Texas 78217	1

Unclassified Security Classification					
DOCUMENT CONT	ROL DATA - R &	s. D			
(Security classification of title, body of abstract and indexing a 1 ORIGINATING ACTIVITY (Corporate author)	annotation must be e	the second s	overall report is classified) CURITY CLASSIFICATION		
			lassified		
Naval Postgraduate School Monterey, California 93940		2b. GROUP			
3 REPORT TITLE					
AN APPROXIMATION TO OFFSET CIRCLE	PROBABILI	TIES			
4 DESCRIPTIVE NOTES (Type of teport and inclusive dates)					
Master's Thesis (March 1971) S. AUTHOR(S) (First name, middle initial, last name)					
George Edmund Wrockloff, III					
Major, United States Army					
6 REPORT DATE	78. TOTAL NO. OF	F PAGES	76. NO. OF REFS		
March 1971	22		6		
88. CONTRACT OF GRANT NO.	98. ORIGINATOR'S	REPORT NUME	ER(S)		
b. PROJECT NO.					
D, PROJECT NO.					
c.	9b. OTKER REPOR this report)	RT NO(S) (Any of	her numbers that may be assigned		
	into reporti				
d. 10. DISTRIBUTION STATEMENT	<u> </u>				
Approved for public release; di	stribution	unlimit	ed		
Approved for public felease, di	SCIEDUCION	UNITINIT C	cu.		
11. SUPPLEMENTARY NOTES	12. SPONSORING	ALLITARY ACTI	VITY		
Naval Postgraduate School Montoroy California 03040					
Monterey, California 93940					
If a weapon system is fired at	t a circula	ar targe	t and the impact		
point is distributed as a bivariate normal random variable, it is not					
particularly difficult to determin					
impact point is at target center. If, for some reason, the expected					
impact point is offset, the problem of determining hit probabilities becomes quite complex. Herein is developed a method to approximate					
such a hit probability.					
· · · ·					
· · · · · · · · · · · · · · · · · · ·					

21

Unclassified

Security	Classification	

14	KEY WORDS				LINK B LINK C		
		ROLE	wτ	ROLE	ΨŤ	ROLE	ΨT
	Offset Circle Probability						
	Approximation to Hit Probability						
	Bessel Function Approximation						
					1		
		•					
		· ·					
DD	FORM \$ A TO (DACK)						
S/N O	FORM 1473 (BACK) 101-807-6821 22 -		Security	Classific	otion		- 31409



